Data Modeling and Least Squares Fitting 2

COS 323

Last time

- Data modeling
- Motivation of least-squares error
- Formulation of *linear* least-squares model:

$$y_i = af(\vec{x}_i) + bg(\vec{x}_i) + ch(\vec{x}_i) + \cdots$$

Given (\vec{x}_i, y_i) , solve for a, b, c, \dots

- Solving using normal equations, pseudoinverse
- Illustrating least-squares with special cases: constant, line
- Weighted least squares
- Evaluating model quality

Nonlinear Least Squares

Some problems can be rewritten to linear

 $y = ae^{bx}$

$$\Rightarrow (\log y) = (\log a) + bx$$

- Fit data points $(x_i, \log y_i)$ to $a^* + bx$, $a = e^{a^*}$
- Big problem: this no longer minimizes squared error!

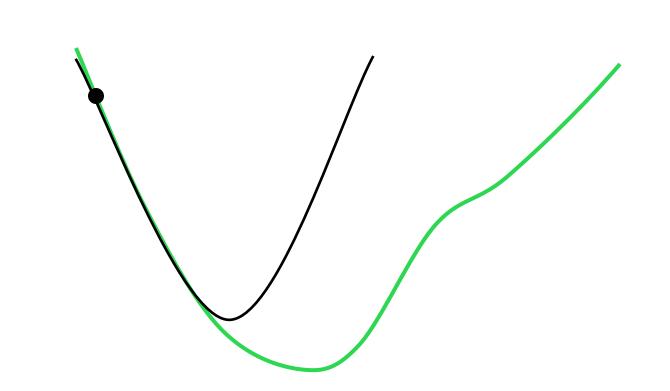
Nonlinear Least Squares

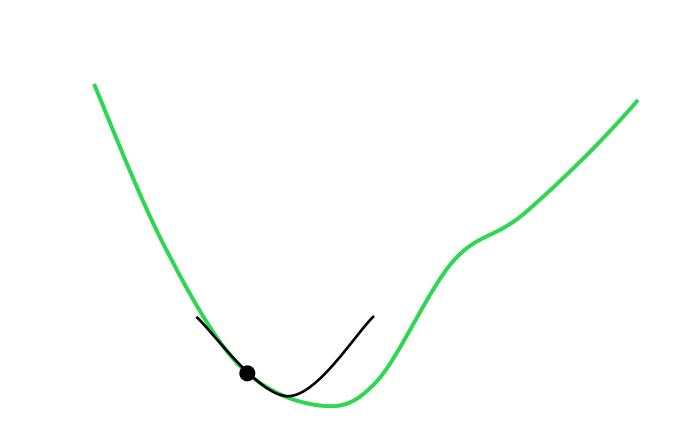
- Can write error function, minimize directly $\chi^{2} = \sum_{i} (y_{i} - f(x_{i}, a, b, ...))^{2}$ Set $\frac{\partial}{\partial a} = 0, \frac{\partial}{\partial b} = 0$, etc.
- For the exponential, no analytic solution for a, b:

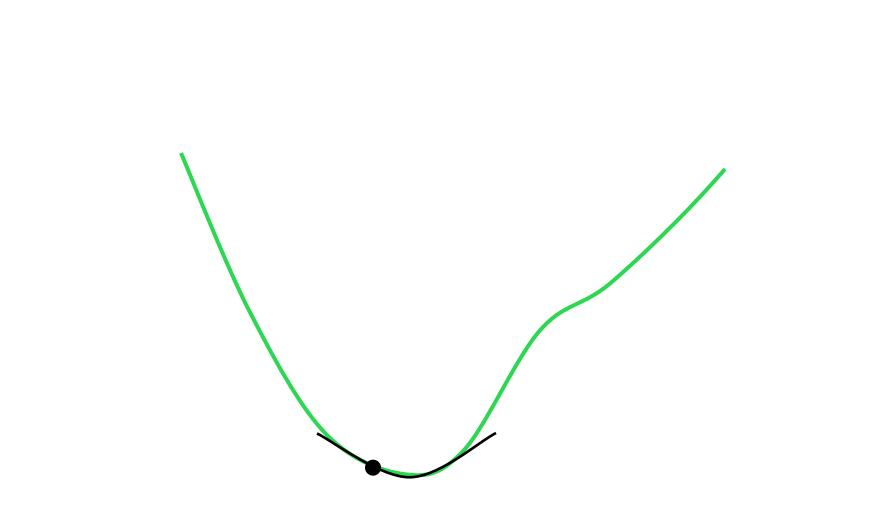
$$\chi^{2} = \sum_{i} \left(y_{i} - ae^{bx_{i}} \right)^{2}$$
$$\frac{\partial}{\partial a} = \sum_{i} -2e^{bx_{i}} \left(y_{i} - ae^{bx_{i}} \right) = 0$$
$$\frac{\partial}{\partial b} = \sum_{i} -2ax_{i}e^{bx_{i}} \left(y_{i} - ae^{bx_{i}} \right) = 0$$

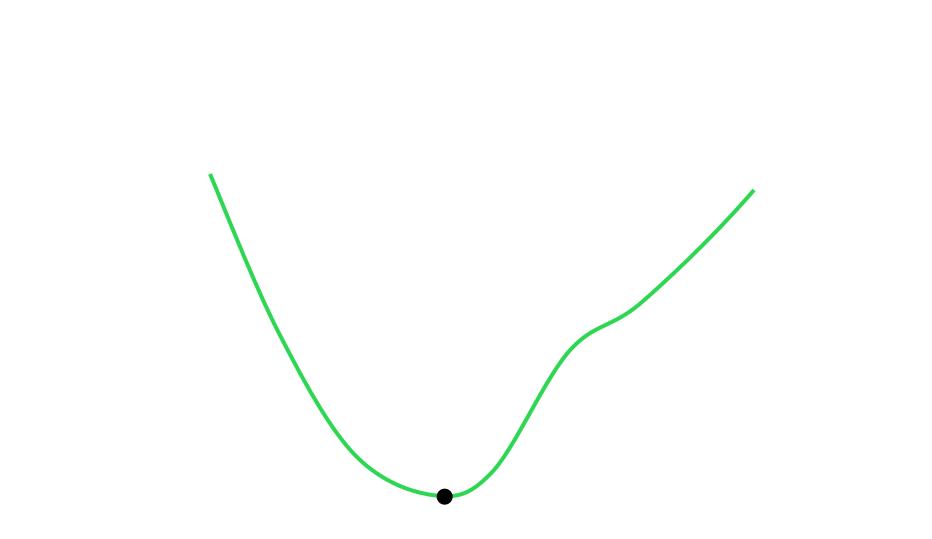
• Apply Newton's method for minimization: - 1-dimensional: $x_{k+1} = x_k - \frac{f'(x_k)}{f''(x_k)}$ - n-dimensional: $\begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i+1} = \begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_i - H^{-1}G$

where H is Hessian (matrix of all 2nd derivatives) and G is gradient (vector of all 1st derivatives)









• Apply Newton's method for minimization: - 1-dimensional: $x_{k+1} = x_k - \frac{f'(x_k)}{f''(x_k)}$ - n-dimensional: $\binom{a}{b}_{i+1} = \binom{a}{b}_i - \mathbf{H}^{-1}\mathbf{G}$

where H is Hessian (matrix of all 2nd derivatives) and G is gradient (vector of all 1st derivatives)

Newton's Method for Least Squares

$$\chi^{2}(a,b,...) = \sum_{i} (y_{i} - f(x_{i},a,b,...))^{2}$$
$$\mathbf{G} = \begin{bmatrix} \frac{\partial(\chi^{2})}{\partial a} \\ \frac{\partial(\chi^{2})}{\partial b} \\ \vdots \end{bmatrix} = \begin{bmatrix} \sum_{i} -2\frac{\partial f}{\partial a}(y_{i} - f(x_{i},a,b,...)) \\ \sum_{i} -2\frac{\partial f}{\partial b}(y_{i} - f(x_{i},a,b,...)) \\ \vdots \end{bmatrix}$$
$$\mathbf{H} = \begin{bmatrix} \frac{\partial^{2}(\chi^{2})}{\partial a^{2}} & \frac{\partial^{2}(\chi^{2})}{\partial a\partial b} & \cdots \\ \frac{\partial^{2}(\chi^{2})}{\partial a\partial b} & \frac{\partial^{2}(\chi^{2})}{\partial b^{2}} & \cdots \\ \vdots \end{bmatrix}$$

• Gradient has 1st derivatives of *f*, Hessian 2nd

Gauss-Newton Iteration

• Consider 1 term of Hessian:

$$\frac{\partial^2(\chi^2)}{\partial a^2} = \frac{\partial}{\partial a} \left(\sum_i -2 \frac{\partial f}{\partial a} (y_i - f(x_i, a, b, \ldots)) \right)$$
$$= -2 \sum_i \frac{\partial^2 f}{\partial a^2} (y_i - f(x_i, a, b, \ldots)) + 2 \sum_i \frac{\partial f}{\partial a} \frac{\partial f}{\partial a}$$

• If close to answer, residual is close to 0, so ignore it \rightarrow eliminates need for 2nd derivatives

Gauss-Newton Iteration

• Consider 1 term of Hessian:

$$\frac{\partial^2(\chi^2)}{\partial a^2} = \frac{\partial}{\partial a} \left(\sum_i -2 \frac{\partial f}{\partial a} (y_i - f(x_i, a, b, \ldots)) \right)$$
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• The Gauss-Newton method approximates $\mathbf{H} \approx 2\mathbf{J}^{\mathrm{T}}\mathbf{J}$ (Only for least-squares!)

Gauss-Newton Iteration

$$\begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i+1} = \begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i} + s_{i}$$

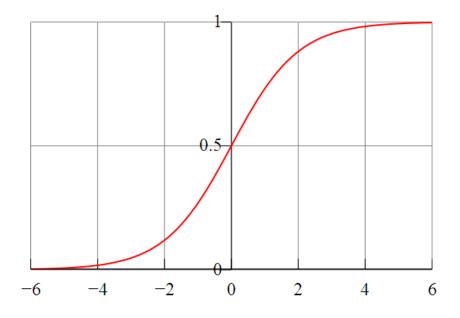
To find Gauss – Newton update $s_i = -\mathbf{H}^{-1}\mathbf{G}$, solve $\mathbf{J}_i^{\mathrm{T}}\mathbf{J}_i s_i = \mathbf{J}_i^{\mathrm{T}} r_i$

$$J = \begin{pmatrix} \frac{\partial f}{\partial a}(x_1) & \frac{\partial f}{\partial b}(x_1) & \dots \\ \frac{\partial f}{\partial a}(x_2) & \frac{\partial f}{\partial b}(x_2) \\ \vdots & \ddots \end{pmatrix}, r = \begin{pmatrix} y_1 - f(x_1, a, b, \cdots) \\ y_2 - f(x_2, a, b, \cdots) \\ \vdots \end{pmatrix}$$

Example: Logistic Regression

 Model probability of an event based on values of explanatory variables, using generalized linear model, logistic function g(z)

$$p(\vec{x}) = g(ax_1 + bx_2 + \cdots)$$
$$g(z) = \frac{1}{1 + e^{-z}}$$





- Assumes positive and negative examples are normally distributed, with different means but same variance
- Applications: predict odds of election victories, sports events, medical outcomes, etc.
- Estimate parameters a, b, ... using Gauss-Newton on individual positive, negative examples
- Handy hint: g'(z) = g(z) (1-g(z))

Gauss-Newton++: The Levenberg-Marquardt Algorithm

Levenberg-Marquardt

- Newton (and Gauss-Newton) work well when close to answer, terribly when far away
- Steepest descent safe when far away
- Levenberg-Marquardt idea: let's do both

$$\begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i+1} = \begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i} - \alpha \mathbf{G} - \beta \begin{pmatrix} \Sigma \frac{\partial f}{\partial a} \frac{\partial f}{\partial a} & \Sigma \frac{\partial f}{\partial a} \frac{\partial f}{\partial b} & \cdots \\ \Sigma \frac{\partial f}{\partial a} \frac{\partial f}{\partial b} & \Sigma \frac{\partial f}{\partial b} \frac{\partial f}{\partial b} & \cdots \end{pmatrix}^{-1} \mathbf{G}$$
Steepest descent Gauss-Newton

Levenberg-Marquardt

- Trade off between constants depending on how far away you are...
- Clever way of doing this:

$$\begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i+1} = \begin{pmatrix} a \\ b \\ \vdots \end{pmatrix}_{i} - \begin{pmatrix} (1+\lambda)\Sigma\frac{\partial f}{\partial a}\frac{\partial f}{\partial a} & \Sigma\frac{\partial f}{\partial a}\frac{\partial f}{\partial b} & \cdots \\ \Sigma\frac{\partial f}{\partial a}\frac{\partial f}{\partial b} & (1+\lambda)\Sigma\frac{\partial f}{\partial b}\frac{\partial f}{\partial b} & \cdots \\ \vdots & \ddots \end{pmatrix}^{-1} \mathbf{G}$$

- If λ is small, mostly like Gauss-Newton
- If λ is big, matrix becomes mostly diagonal, behaves like steepest descent

Levenberg-Marquardt

- Final bit of cleverness: adjust λ depending on how well we're doing
 - Start with some λ , e.g. 0.001
 - If last iteration decreased error, accept the step and decrease λ to $\lambda/10$
 - If last iteration *increased* error, *reject* the step and *increase* λ to 10λ
- Result: fairly stable algorithm, not too painful (no 2nd derivatives), used a lot

Dealing with Outliers

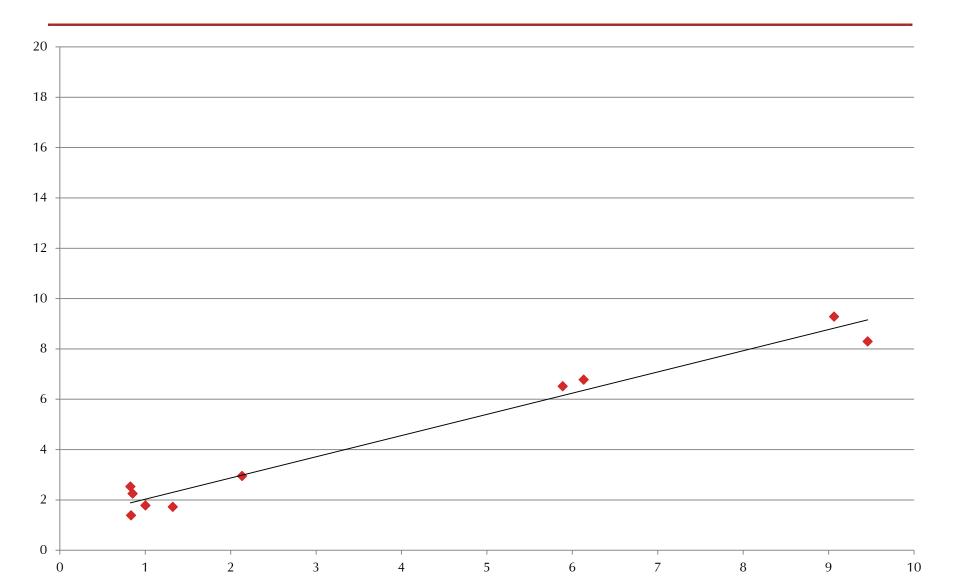
Outliers

- A lot of derivations assume Gaussian distribution for errors
- Unfortunately, nature (and experimenters) sometimes don't cooperate

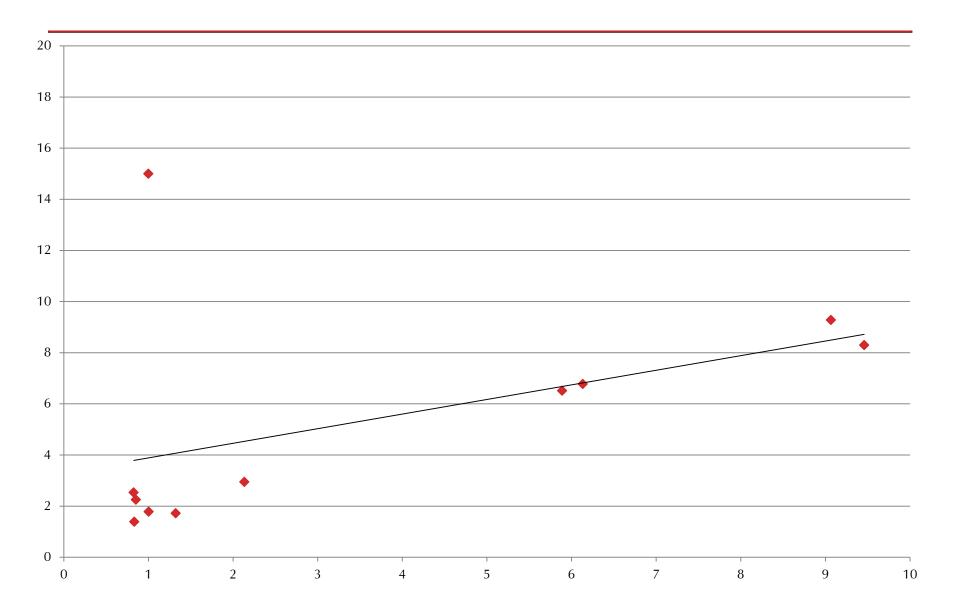
Non-Gaussian

- Outliers: points with extremely low probability of occurrence (according to Gaussian statistics)
- Can have strong influence on least squares

Example: without outlier



Example: with outlier



Robust Estimation

- Goal: develop parameter estimation methods insensitive to *small* numbers of *large* errors
- General approach: try to give large deviations less weight
- e.g., Median is a robust measure, mean is not
- M-estimators: minimize some function other than square of y – f(x,a,b,...)

Least Absolute Value Fitting

- Minimize $\sum_{i} |y_i f(x_i, a, b, ...)|$ instead of $\sum_{i}^{i} (y_i - f(x_i, a, b, ...))^2$
- Points far away from trend get comparatively less influence

Example: Constant

- For constant function y = a, minimizing $\Sigma(y-a)^2$ gave a = mean
- Minimizing $\Sigma |y-a|$ gives a = median

Least Squares vs. Least Absolute Deviations

• LS:

- Not robust
- Stable, unique solution
- Solve with normal equations, Gauss-Newton, etc.
- LAD
 - Robust
 - Unstable, not necessarily unique
 - Nasty function (discontinuous derivative):
 requires iterative solution method (e.g. simplex)

Iteratively Reweighted Least Squares

 Sometimes-used approximation: convert to iteratively weighted least squares

$$\sum_{i} |y_{i} - f(x_{i}, a, b, ...)|$$

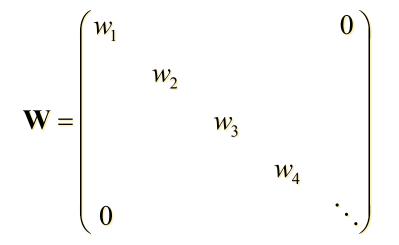
$$= \sum_{i} \frac{1}{|y_{i} - f(x_{i}, a, b, ...)|} (y_{i} - f(x_{i}, a, b, ...))^{2}$$

$$= \sum_{i} w_{i} (y_{i} - f(x_{i}, a, b, ...))^{2}$$

with w_i based on previous iteration

Review: Weighted Least Squares

• Define weight matrix W as



• Then solve weighted least squares via

 $\mathbf{A}^{\mathrm{T}}\mathbf{W}\mathbf{A}\,x = \mathbf{A}^{\mathrm{T}}\mathbf{W}\,b$

M-Estimators

Different options for weights

- Give even less weight to outliers

$$w_{i} = \frac{1}{|y_{i} - f(x_{i}, a, b, ...)|} \qquad L_{1}$$

$$w_{i} = \frac{1}{\varepsilon + |y_{i} - f(x_{i}, a, b, ...)|} \qquad "Fair"$$

$$w_{i} = \frac{1}{\varepsilon + (y_{i} - f(x_{i}, a, b, ...))^{2}} \qquad Cauchy / Lorentzian$$

$$w_{i} = e^{-k(y_{i} - f(x_{i}, a, b, ...))^{2}} \qquad Welsch$$

Iteratively Reweighted Least Squares

- Danger! This is not guaranteed to converge to the right answer!
 - Needs good starting point, which is available if initial least squares estimator is reasonable
 - In general, works OK if few outliers, not too far off

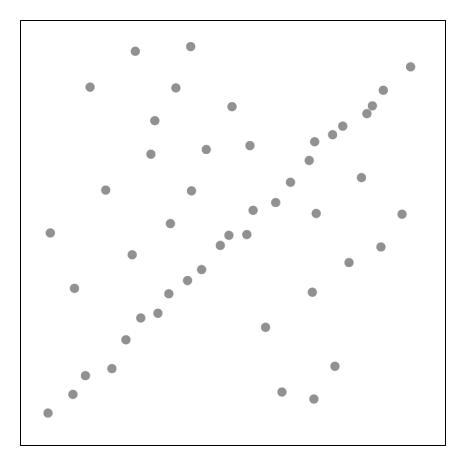
Outlier Detection and Rejection

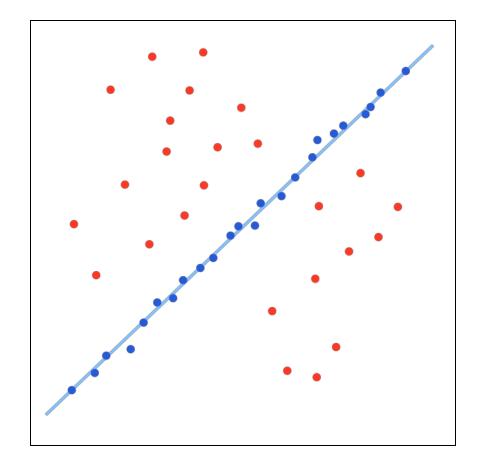
- Special case of IRWLS: set weight = 0 if outlier,
 1 otherwise
- Detecting outliers: $(y_i f(x_i))^2 >$ threshold
 - One choice: multiple of mean squared difference
 - Better choice: multiple of *median* squared difference
 - Can iterate...
 - As before, not guaranteed to do anything reasonable, tends to work OK if only a few outliers

RANSAC

- RANdom SAmple Consensus: desgined for bad data (in best case, up to 50% outliers)
- Take many *minimal* random subsets of data
 Compute fit for each sample
 - See how many points agree: $(y_i f(x_i))^2 < \text{threshold}$
 - Threshold user-specified or estimated from more trials
- At end, use fit that agreed with most points
 Can do one final least squares with all inliers

RANSAC





Least Squares in Practice

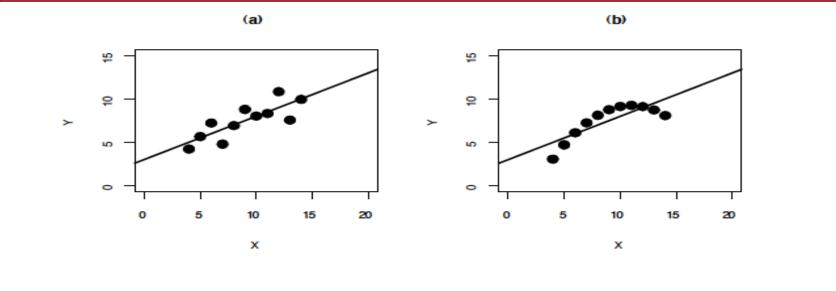
Least Squares in Practice

- More data is better $\sigma^2 = \frac{\chi^2}{n-m}C$ - uncertainty in estimated parameters goes down slowly: like 1/sqrt(# samples)
- Good correlation doesn't mean a model is good
 use visualizations and reasoning, too.

Anscombe's Quartet

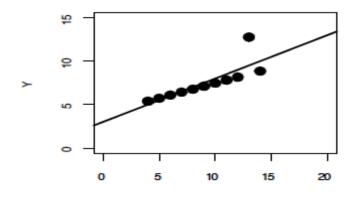
Dataset 1		Dataset 2		Dataset 3		Dataset 4		
X	у	X	у	X	у	x	у	
10	8.04	10	9.14	10	7.46	8	6.58	
8	6.95	8	8.14	8	6.77	8	5.76	
13	7.58	13	8.74	13	12.74	8	7.71	
9	8.81	9	8.77	9	7.11	8	8.84	y = 3.0 + 0.5x
11	8.33	11	9.26	11	7.81	8	8.47	r = 0.82
14	9.96	14	8.10	14	8.84	8	7.04	
6	7.24	6	6.13	6	6.08	8	5.25	
4	4.26	4	3.10	4	5.39	19	12.50	
12	10.84	12	9.13	12	8.15	8	5.56	
7	4.82	7	7.26	7	6.42	8	7.91	
5	5.68	5	4.74	5	5.73	8	6.89	

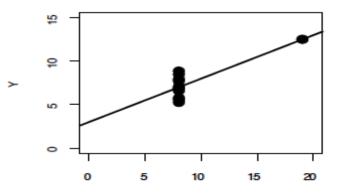
Anscombe's Quartet



(c)

(d)



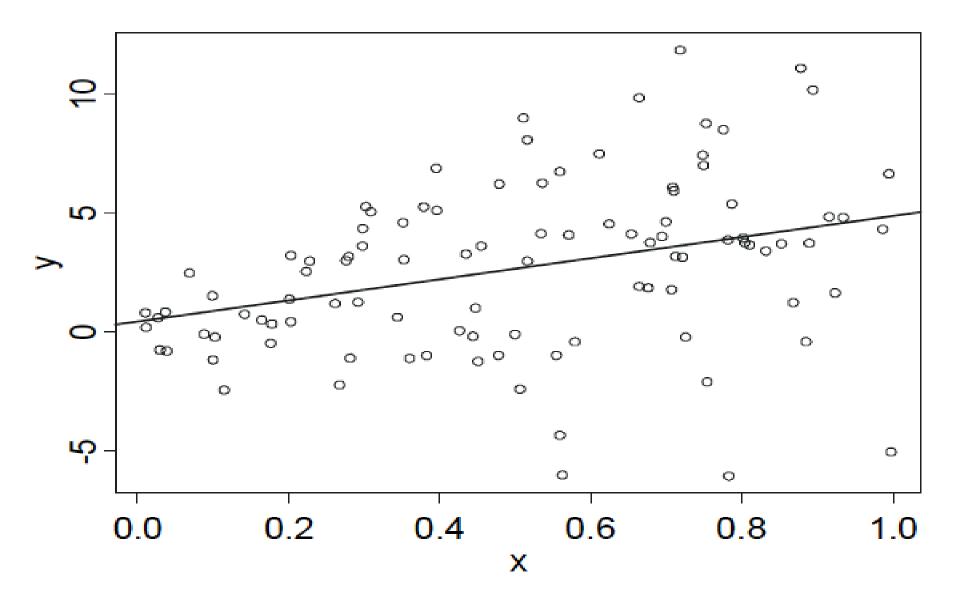


x

х

Least Squares in Practice

- More data is better
- Good correlation doesn't mean a model is good
- Many circumstances call for (slightly) more sophisticated models than least squares
 - Generalized linear models, regularized models (e.g., LASSO), PCA, ...

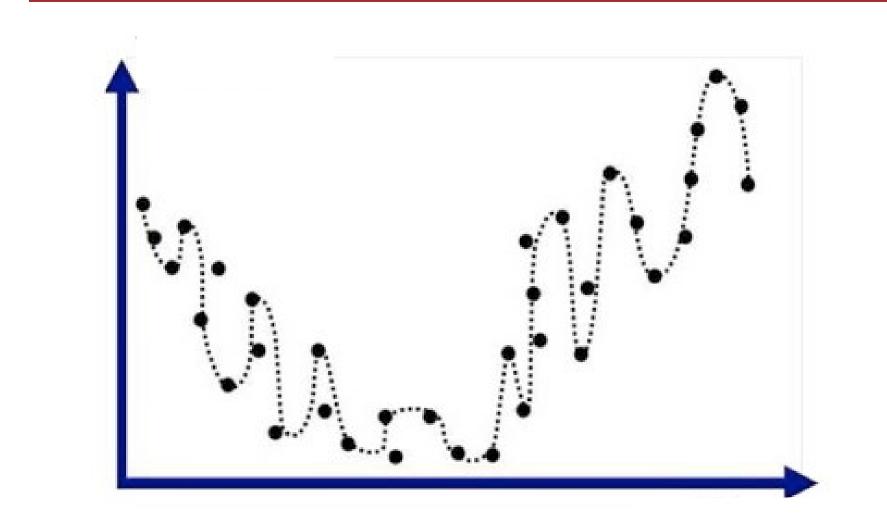


Residuals depend on x (heteroscedastic): Assumptions of linear least squares not met

Least Squares in Practice

- More data is better
- Good correlation **doesn't mean a model is good**
- Many circumstances call for (slightly) more sophisticated models than linear LS
- Sometimes a model's fit can be too good ("overfitting")
 - more parameters may make it easier to overfit

Overfitting



Least Squares in Practice

- More data is better
- Good correlation **doesn't mean a model is good**
- Many circumstances call for (slightly) more sophisticated models than linear LS
- Sometimes a model's fit can be **too good**
- All of these minimize "vertical" squared distance
 Square, vertical distance not always appropriate