

MATHEMATICAL METHODS IN THEORETICAL CS
LECTURE 6: RAZBOROV DISJOINTNESS LOWER BOUND,
FORSTER'S THEOREM

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Summary: In this lecture, we show two results dealing with lower bounds in communication complexity. The first lower bound is an $\Omega(n)$ lower bound on the distributional complexity of *Disjointness* due to [3, 8]. Here we will present the simplified proof presented in [8]. In the second part, we will show how to obtain lower bounds on the unbounded error probabilistic communication complexity by Forster's method [2] of lower-bounding the sign rank of the corresponding matrix by showing that it has a small spectral norm.

1. DISJOINTNESS LOWER BOUND

In the first part of the lecture, we will try to lower bound the ϵ -error probabilistic communication complexity of a predicate A , $C_\epsilon(A)$ by using the concept of the Distributional complexity.

The ϵ -error distributional communication complexity $D_\epsilon^\mu(A)$ is the minimum cost (w.r.t. number of communications, measured in say, bits) deterministic protocol P such that

$$(1) \quad \Pr_\mu[P(X, Y) = A(X, Y)] \geq 1 - \epsilon$$

where μ is a given probability distribution over the inputs X and Y (each n bits long) to the two parties. Yao showed that $D_\epsilon^\mu(A)$ can be used to lower bound $C_\epsilon(A)$ as $C_\epsilon(A) \geq \frac{1}{2} D_{2\epsilon}^\mu(A)$ [5].

We will now prove a $\Omega(n)$ lower bound on probabilistic communication complexity of the Disjointness function by constructing a μ for which $D_\epsilon^\mu(DISJ_n) = \Omega(n)$ where $DISJ_n$ is the *Disjointness* predicate, where the inputs X and Y are $\in \{0, 1\}^n$ each representing a subset of $\{1, 2, \dots, n\}$ (represented as $[n]$).

Theorem 1.1. $\exists \mu$ such that $D_\epsilon^\mu(DISJ_n) \geq \Omega(n)$ where $\epsilon < \frac{1}{100}$, where

$$(2) \quad DISJ(X, Y) = \begin{cases} 1 & \text{if } X \cap Y = \emptyset \\ 0 & \text{otherwise} \end{cases}$$

Proof. We first give the probability distribution μ over the inputs. Let $n = 4m - 1$ and let $T = (T_x, T_y, \{i\})$ be an arbitrary partition of $[n]$ such that $|T_x| = |T_y| = 2m - 1$. Now, the input to the first party, an m -element set X ($|X| = m$) is chosen uniformly at random from $T_x \cup \{i\}$, and the m -element set Y is similarly chosen uniformly at random from $T_y \cup \{i\}$. Let $X_0(Y_0)$ correspond to an input X (resp. Y) such that $i \notin X$ (resp. Y), and X_1 (resp. Y_1) correspond to an input such that $i \in X$ (resp. Y). Hence, in the distribution defined, we get inputs of the types (X_0, Y_0) , (X_0, Y_1) , (X_1, Y_0) and (X_1, Y_1) with an equal probability of $\frac{1}{4}$ each.

Let A be the set of inputs of the type (X_1, Y_1) ($DISJ$ value 1) and B be the other inputs with non-zero weight ($DISJ$ value 0). The proof of the theorem follows from Lemma 1.5 as follows.

Let $D_\epsilon^\mu = k$. Let R_1, R_2, \dots, R_t (with $t \leq 2^k$) be the (almost) monochromatic rectangles which have the function value 1.

$$\begin{aligned} \mu(B \cap \bigcup_{i=1}^t R_i) &\leq \epsilon \\ \mu(B \cap \bigcup_{i=1}^t R_i) &= \sum_{i=1}^t \mu(R_i \cap B) \\ &\geq \sum_{i=1}^t \alpha \mu(A \cap R_i) - 2^{-\delta n} \text{ (from Lemma 1.5)} \\ &\geq \alpha \left(\frac{3}{4} - \epsilon\right) - t 2^{-\delta n} \end{aligned}$$

Hence by choosing a small enough ϵ , we get $k = \Omega(n)$. \square

The lemma 1.5 shows that in any rectangle (of inputs) which is not too small (weight more than $\frac{2^{-\delta n}}{\alpha}$), $R = \mathbb{X} \times \mathbb{Y}$, the number of inputs which return a (disjointness) function value 0 is not too small, hence showing that no rectangle (which is not too small) is close to being monochromatic. However, we show another lemma with a similar flavour first.

Lemma 1.2. For any $R = \mathbb{X} \times \mathbb{Y}$, where $\mathbb{X}, \mathbb{Y} \subseteq 2^{[n]}$,

$$(3) \quad P[(x_1, y_1) \in R] \geq \alpha Pr[(x_0, y_0) \in R] - 2^{-\Omega(n)}$$

Proof. Let $t = (t_X, t_Y, \{i\})$ be a partition of $[n]$. For ease of notation, we define the following terms:

$$\begin{aligned} (4) \quad p(t) &= Pr(X \in \mathbb{X} | T = t) \\ (5) \quad p_0(t) &= Pr(X_0 \in \mathbb{X} | T = t) \\ (6) \quad p_1(t) &= Pr(X_1 \in \mathbb{X} | T = t) \\ (7) \end{aligned}$$

Further, we call a partition $t = (t_X, t_Y, \{i\})$ *X-bad* if

$$(8) \quad p_1(t) < \frac{1}{3} p_0(t) - 2^{-\epsilon n}$$

Similarly we define the terms $q(t), q_0(t), q_1(t)$ and *Y-bad* for the inputs to the second party Y . We also call a partition t bad iff it is either *X-bad* or *Y-bad*. To prove the lemma, we would like to show that most partitions are not bad (claim 1.3).

We first observe that by fixing the partition $T = (T_X, T_Y, \{i\})$, the two quantities $Pr(X_1 \in R)$ and $Pr(Y_1 \in R)$ becomes independent¹. Hence

$$(9) \quad Pr[(X_\lambda, Y_\lambda) \in R] = \mathbb{E}_t[p_\lambda(t)q_\lambda(t)] \text{ for } \lambda \in \{0, 1\}$$

¹This method of fixing t to make the two events independent in order to get a convex combination has also been used subsequently in the proof of the Parallel Repetition theorem by Raz [7].

We also observe that

$$(10) \quad p(t) = \frac{1}{2}(p_0(t) + p_1(t))$$

Finally, we note that fixing t_Y , fixes $p(t)$ and $q_0(t)$ and fixing t_X , also fixes $q(t)$ and $p_0(t)$.

Now, we will proceed to show the lemma component-wise.

Claim 1.3. *For every set $t_Y \subseteq [n]$ such that $|t_Y| = 2m - 1$,*

$$(11) \quad \Pr(T \text{ is } X\text{-bad} | T_Y = t_Y) < \frac{1}{5}$$

Proof. Given t_Y , $X \leftarrow_{\mathbb{R}} [n] \setminus t_Y$ (with $|X| = m$), and hence $\Pr(T \text{ is } X\text{-bad} | T_Y = t_Y)$ becomes fixed too.

If $p(t) < 2^{-\epsilon n}$, from equation 10, $p_0(t) \leq 2p(t)$. Hence, if t is X -bad, $p_1(t) < \frac{1}{3}p_0(t) - 2^{-\epsilon n} < -\frac{1}{3}2^{-\epsilon n} < 0$, which is impossible.

Consider the case when $p(t) \geq 2^{-\epsilon n}$. Let $\Gamma = \mathbb{X} \cap \{X | X \subseteq [n] \text{ s.t. } |X| = m\}$.

$p(t) = \frac{|\Gamma|}{\binom{[n]}{m}}$. Also, if $s \leftarrow_{\mathbb{R}} \Gamma$, then,

$$p_0(t) = 2p(t)\Pr(i \in s)$$

This follows because,

$$\begin{aligned} p_0(t) &= \Pr(X \in \mathbb{X} | t = t, i \notin X) \\ &= 2\Pr(X \in \mathbb{X}, i \notin X | T = t) \\ &= 2\Pr(i \notin X | X \in \mathbb{X}, T = t)\Pr(X \in \mathbb{X} | T = t) \\ &= 2p(t)\Pr[i \notin s] \end{aligned}$$

Similarly,

$$p_1(t) = 2p(t)\Pr(i \notin s)$$

Now, if partition t is X -bad, we have from equations 8,1,1

$$(12) \quad \Pr[i \in s] < \frac{1}{3}\Pr[i \notin s] - \frac{2^{-\epsilon n}}{2p(t)}$$

Since $p(t) \geq 2^{-\epsilon n}$, $\Pr[i \in s] < \frac{1}{3}\Pr[i \notin s]$. Hence,

$$(13) \quad \Pr(i \in s) < \frac{1}{4}$$

Let $\{i_1, i_2, \dots, i_{2m}\} = [n] - T_y$ and let $\vec{s} = (s_1, s_2, \dots, s_{2m})$ where the s_j indicates whether $i_j \in s$. We now show that claim by using an entropy argument on the possible choice of vectors \vec{s} . Let the claim 1.3 not hold, in which case

$\Pr[T = (T_X, T_Y, \{i\}) \text{ is } X\text{-bad} | T_Y = T_y] \geq \frac{1}{5}$. Then, calculating the entropy we get

$$\begin{aligned} H(s) &\geq m(2 - 4\epsilon - o(1)) \\ H(s) &\leq \sum_i 1^{2m} H(s_i) \\ &\leq \frac{8m}{5} + \frac{2m}{5} H\left(\frac{1}{4}\right) \\ &\leq 1.93m \end{aligned}$$

which is a contradiction if we choose a small enough ϵ . Hence the claim is true. \square

Let $Bad(T)$ denote the indicator of the event $T = (T_X, T_Y, \{i\})$ being Bad , and let $Bad_X(T), Bad_Y(T)$ be the indicator of the events T being X - Bad and Y - Bad respectively. In the next claim, we prove that contribution of non- Bad partitions T to the RHS of Lemma 1.2 is insignificant.

Claim 1.4.

$$\mathbb{E}_T[p_0(T)q_0(T)Bad(T)] \leq \frac{4}{5}\mathbb{E}_T[p_0(T)q_0(T)]$$

Proof. Since $Bad(T) \leq Bad_X(T) + Bad_Y(T)$, and by symmetry, it suffices to show that

$$\mathbb{E}_T[p_0(T)q_0(T)Bad_X(T)] \leq \frac{2}{5}\mathbb{E}_T[p_0(T)q_0(T)]$$

Fixing T_Y also fixes $p(T), q_0(T)$. Hence,

$$\begin{aligned} \mathbb{E}[p_0(T)q_0(T)Bad_X(T)|T_Y = t_Y] &= 2pq_0\mathbb{E}_T[Bad_X|T_Y = t_Y] \\ &\leq \frac{2}{5}pq_0 \text{ from claim 1.3} \\ &\leq \frac{2}{5}\mathbb{E}_T[q_0(T)|T_Y = t_Y] \\ &\leq \frac{2}{5}\mathbb{E}_T[p_0(T)q_0(T)|T_Y = t_Y] \end{aligned}$$

□

Now, to complete the proof of the lemma (1.2),

$$\begin{aligned} P[(X_1, Y_1) \in R] &= \mathbb{E}_T[p_1(T)q_1(T)] \\ &\geq \mathbb{E}_T[p_1(T)q_1(T)(1 - Bad(T))] \\ &\geq \mathbb{E}_T\left[\left(\frac{1}{3}p_0(T) - 2^{-\epsilon n}\right)\left(\frac{1}{3}q_0(T) - 2^{-\epsilon n}\right)(1 - Bad(T))\right] \text{ (from eq 8)} \\ &\geq \alpha\mathbb{E}_T[p_0(T)q_0(T)] - 2^{-\Omega(n)} \\ &\geq \alpha Pr[(X_0, Y_0) \in R] - 2^{-\Omega(n)} \end{aligned}$$

□

Now, we prove the lemma required in the proof of the theorem.

Lemma 1.5. For any $R = \mathbb{X} \times \mathbb{Y}$, where $\mathbb{X}, \mathbb{Y} \subseteq 2^{[n]}$,

$$(14) \quad \mu(B \cap R) \geq \alpha\mu(A \cap R) - 2^{-\delta n}$$

for some constants $k, \delta > 0$.

Proof. This lemma follows from the previous lemma (1.2) by just observing that

$$\begin{aligned} \mu(B \cap R) &= \frac{1}{4}\mathbb{E}_T[p_1(T)q_1(T)] \\ \mu(A \cap R) &= \frac{3}{4}\mathbb{E}_T[p_0(T)q_0(T)] \end{aligned}$$

We prove just the first equation. The second equation follows along similar lines.

$$\begin{aligned}
 \mu(B \cap R) &= \mu(B)\mu(R|B) \\
 &= \frac{1}{4} \sum T \Pr[T] \Pr[X \in \mathbb{X}|T = t, i \in X] \Pr[Y \in \mathbb{Y}|T = t, i \in Y] \\
 &= \frac{1}{4} \mathbb{E}_T [p_1(T)q_1(T)]
 \end{aligned}$$

□

2. FORSTER'S LOWER BOUND

In this section, we present Forster's lower bound for the *Unbounded Error Probabilistic Communication Complexity* by showing a lower bound on the *sign rank* of the corresponding communication matrix, if it has a low spectral norm. In [6], it was shown that the communication complexity (C_f) of a distributed function f is closely related to *sign rank* (say k) (2.1) of the communication matrix as

$$(15) \quad \lceil \log_2 k \rceil \leq C_f \leq \lceil \log_2 k \rceil + 1$$

For a distributed function $f : \{1, 2, \dots, n\}^2 \rightarrow \{-1, 1\}$ represented by the matrix $M(f) \in \{-1, 1\}^{n \times n}$, we define the sign rank of the corresponding matrix.

Definition 2.1. For a matrix $M \in \{-1, 1\}^{n \times n}$, we say that the *signrank*(M) $\leq k$ iff there exists $A \in \mathbb{R}^{n \times n}$ of rank $\leq k$ such that $M_{i,j} = \text{sign}(A_{i,j})$ (where $\text{sign}(x)$ is the usual sign function defined on \mathbb{R}). Equivalently, there exists $X = \{x_1, x_2, \dots, x_n\}, Y = \{y_1, y_2, \dots, y_n\} \subseteq \mathbb{R}^k$ such that $M_{i,j} = \text{sign}(\langle x_i, y_j \rangle)$, where $\langle a, b \rangle$ represents the inner product of vectors a and b .

We now present and prove Forster's theorem [2]

Theorem 2.2. *signrank*(M) $\geq \frac{n}{\|M\|}$ where $\|A\|$ represents the spectral norm of the matrix A .

Note that this theorem along with theorem 15 implies that for any distributed function $f : \{0, 1\}^m \times \{0, 1\}^m \rightarrow \{0, 1\}$, the communication complexity

$$C_f \geq m - \log_2 \|M_f\|$$

where M_f represents the corresponding communication matrix. For the sake of notation, for vectors $x \in X$ and $y \in Y$, we refer to the corresponding entry in M by $M_{x,y}$.

Proof. The proof follows in two steps. The first is in establishing a connection between a relaxation of the Discrepancy($\text{disc}(M)$) and the spectral norm $\|M\|$. The second part of the proof lower bounds disc by using Lemma 2.3, which forms the crux of the proof. We now state the lemma (proved in the special note by David Steurer)

Lemma 2.3. For every $X \subseteq \mathbb{R}^k$ such that $|X| = n$, such that all subsets of X with at most k elements are linearly independent, there exists a linear map $A \in GL(k)$ such that

$$\sum_{x \in X} \frac{1}{\|Ax\|^2} (Ax)(Ax)^T = \frac{n}{k} I_k$$

We first define $disc(M)$.

$$disc(M) = \max_{\substack{X=\{x|||x||=1\} \\ Y=\{y|||y||=1\} \\ |Y|=|X|=n}} \sum_{x \in X, y \in Y} M_{x,y} \langle x, y \rangle$$

This is related to within a constant factor of the earlier definition of discrepancy by the Groethendieck's inequality ([1]). Now, we establish the following relation $disc \leq n||M||$.

$$(16) \quad ||M|| = \max_{||u||=1, ||v||=1} \langle u, Mv \rangle$$

We now show that

$$(17) \quad ||M|| = \max_{\substack{\sum_x ||x||^2=1 \\ \sum_y ||y||^2=1}} M_{x,y} \langle x, y \rangle$$

Clearly, $LHS \leq RHS$. We now show the other direction by apply Cauchy-Schwartz inequality twice.

$$\begin{aligned} \max_{\substack{\sum_x ||x||^2=1 \\ \sum_y ||y||^2=1}} M_{x,y} \langle x, y \rangle &\leq \sum_i \sum_{\substack{\sum_x ||x||^2=1 \\ \sum_y ||y||^2=1}} M_{x,y} x_i y_i \\ &\leq \sum_i ||M|| \sqrt{(\sum_x x_i^2)(\sum_y y_i^2)} \\ &\leq ||M|| \sum_i \sqrt{\sum_x x_i^2} \sqrt{\sum_y y_i^2} \\ &\leq ||M|| \sqrt{\sum_i \sum_x x_i^2} \sqrt{\sum_i \sum_y y_i^2} \\ &\leq ||M|| \end{aligned}$$

Finally, we proceed to prove the theorem. M has sign rank k , and assume that the corresponding unit vectors are $x \in X$ and $y \in Y$.

$$\begin{aligned} \sum_{x,y} M_{x,y} \langle x, y \rangle &= \sum_{x,y} | \langle x, y \rangle | \\ &\geq \sum_{x,y} \langle x, y \rangle^2 \\ &= \sum_y Y^T (\sum_x x x^T) Y \\ &= \frac{n^2}{k} \text{ from Lemma 2.3} \\ \text{Hence } n||M|| &\geq disc \\ &\geq \frac{n^2}{k} \end{aligned}$$

Thus, $||M|| \geq \frac{n}{k}$.

□

In particular, this result by Forster also resolved a long-standing conjecture of [6, 4] in showing that the unbounding error probabilistic communication complexity of the distributed function given by the Hadamard matrix is linear ($\geq \frac{n}{2}$).

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