Model-Based Classification

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Probability models

- A probability model is a joint distribution of a set of observations.
- Often, a model is indexed by a *parameter*. Each value of the parameter gives a different distribution of the data.
 - The parameter of a Bernoulli is the probability of heads.
 - The parameters of a Gaussian are its mean and variance.
- Many models (but not all) assume the data are independent and identically distributed.
- For a boring example, consider N coin flips, each of which has heads with probability π ,

$$p(x_1,...,x_N | \pi) = \prod_{n=1}^{N} p(x_i | \pi).$$
(1)

Each term is a Bernoulli,

$$p(x_n | \pi) = \pi^{1(x_n = h)} (1 - \pi)^{1(x_n = t)}$$
(2)

- Suppose we flip a coin *N* times and record the outcomes.
- Further suppose that *we think* that the probability of heads is π . (This is distinct from whatever the probability of heads "really" is.)
- Given π , the probability of an observed sequence is

$$p(x_1, \dots, x_N | \pi) = \prod_{n=1}^N \pi^{1[x_n = h]} (1 - \pi)^{1[x_n = t]}$$
(3)

- As a function of π , the probability of a data set is the *likelihood function*.
- Taking logs, this is the log likelihood function.

$$\mathscr{L}(\pi) = \sum_{n=1}^{N} \mathbb{1}[x_n = \mathbf{h}] \log \pi + \mathbb{1}[x_n = \mathbf{t}] \log(1 - \pi)$$
(4)

- The *maximum likelihood estimate* is the value of the parameter that maximizes the log likelihood (equivalently, the likelihood).
- In the Bernoulli example, it is the proportion of heads.

$$\hat{\pi} = \frac{1}{N} \sum_{n=1}^{N} \mathbb{1}[x_n = H]$$
(5)

• In a Gaussian, it is the empirical mean

$$\hat{\mu} = \frac{1}{N} \sum_{n=1}^{N} x_n \tag{6}$$

and empirical variance

$$\hat{\sigma}^2 = \frac{1}{N} \sum_{n=1}^{N} (x_n - \hat{\mu})^2 \tag{7}$$

- In a sense, this is the value that best explains our observations.
- Why is the MLE good?
 - The MLE is *consistent*.
 - Flip a coin *N* times with true bias π^* .
 - Estimate the parameter from $x_1, \ldots x_N$ with the MLE $\hat{\pi}$.
 - Then,

$$\lim_{N\to\infty}\hat{\pi}=\pi^*$$

- This is a good thing. It lets many statisticians sleep at night.
- (R demonstration: Bernoulli data sets)
- (Slides: Modeling approval ratings with a Gaussian)

Graphical models

- Represents a joint distribution of random variables; used to show models. (Also called "Bayesian network")
- Semantics:
 - Nodes are RVs; Edges denote possible dependence
 - Shaded nodes are observed; unshaded nodes are hidden
 - GMs with shaded nodes represent posterior distributions.
- Each graphical model is a family of distributions.
- Connects computing about models to graph structure (COS513)
- Connects independence assumptions to graph structure (COS513)
- Here we'll use it as a schematic for factored joint distributions. (Show the classification graphical model.)
- Return briefly to the Aliens/Watch example
 - It's true that many members of this family do not have $X \perp Y | Z$.
 - The class discussion revealed conditions where *Z* is dependent on *X* and *Y*, but they are still conditionally independent. (That is a subfamily of this family.)
 - I conjectured that for conditional independence between *X* and *Y*, one of the dependencies of *Z* had to be broken. However, I couldn't prove it.

Classification set-up

- In classification we observe two sets of data. One set contains data ("features") labeled with a category. The other set contains unlabeled features.
- The idea is to fit a model of how the label relates to the features. Then given new unlabeled data, predict the category. This is *supervised learning*.
- More formally,

- The fully observed data (also called the "training data") are $\{x_i, c_i\}_{i=1}^n$, where c_i is in one of k categories and the features x_i are a vector of values.
- The partially observed data are x_{new} . Our goal is to predict y_{new} .
- Here are some examples. In each, what are the features? What are the labels?
 - Classify images into semantic categories
 - Classify news articles into section of the newspaper
 - Classify genetic code as entron or intron
 - Classify radar blips as friendly or unfriendly
 - Classify credit cards as stolen or not stolen
 - Others?

Basic idea

- Classification with generative models links statistical modeling to classification problems.
- We can model many kinds of data (features *x*) with appropriate probability distributions.
 - Continuous features can be modeled with a Gaussian
 - Binary features can be modeled with a Bernoulli
 - Positive features can be modeled with a Gamma
 - Etc.
- Recall that our training set is a collection of labeled feature vectors (x_i, c_i) . The idea is to fit a distribution of features *conditional* on the class label.
 - A different Gaussian for each class label
 - A different Bernoulli for each class label
 - A different Gamma distribution for each class label
 - Etc.
- To classify a new feature vector x we compute the conditional distribution of the class label given the features p(c|x).
- For example:

- Suppose images are represented with continuous value image features such as color intensities, texture features, and others.
- Each training image is labeled with one category, such as "outdoor", "indoor", "sports". (This data set exists. It's called CalTech 101.)
- To build our classifier, we fit a Gaussian distribution to each feature conditional on the class label. For example, we would find a distribution of color for indoor scenes, for outdoor scenes, for portraits, and each other category.¹
- To classify a new image, we would consider each label and look at the probability of its features given that label. (It's a little more complicated than this—see below—and this process will emerge naturally from Bayes rule.)

Modeling assumptions

- Data
 - The training data are $\mathcal{D} = \{x_i, c_i\}_{i=1}^n$
 - The data to predict are $\{x_{new}\}$.
- The graphical models for fitting and prediction are the following,



- Some details
 - Small boxes are parameters. Unshaded are fit; shaded are fixed.
 - The joint distribution of a feature vector and label is

$$p(x, c | \pi, \theta_{1:k}) = p(c | \pi) p(x | c, \theta_{1:k})$$
(8)

- The parameters $\theta_{1:k}$ are the conditional distributions of the features. The parameter π is the probability of seeing each class.

¹Actually, we could fit a conditional distribution for the whole vector using a multivariate Gaussian. But, for now, let's assume we fit one distribution for each feature.

- The second term $p(x | c, \theta_{1:k})$ selects the right class. (More below.)
- How do we "select" the right class in $p(x | c, \theta_{1:k})$ and $p(c | \pi)$?
 - The class label *c* is represented as a *k*-vector with a single one. For example, a data point in the fourth class has $c = \langle 0, 0, 0, 1, 0, 0, 0 \rangle$.
 - The second term is

$$p(x | c, \theta_{1:k}) = \prod_{j=1}^{k} p(x | \theta_j)^{c^j}.$$
(9)

(Confirm that this equals the intended probability.)

- The distribution in π works the same way. The parameter π is a *k*-vector of probabilities that sum to one. The probability of selecting a particular class label is

$$p(c \mid \pi) = \prod_{j=1}^{k} \pi_j^{c^j}.$$
 (10)

Note: the space of positive vectors that sum to one is called the *simplex*.

- As we'll see, this representation helps in fitting the model.

- In *fitting*, we want to find the class parameters θ_k and class proportions π from a data set of labeled observations.
- In *prediction*, we use fitted parameters to predict the label of an unlabeled data point. We need to compute *p*(*c*|*x*).
- Example #1 : Gaussian classification
 - Let's say the data are continuous.
 - To be concrete, suppose each x_i is an image, with a vector real-valued image features measured on it. (E.g., these might be texture, color histogram, etc.)
 - Suppose each class is described by a Gaussian, where $\theta_k = \mu_k$ and $x | \theta_k \sim \mathcal{N}(\mu_k, \sigma^2)$. (Each shares the same variance, for simplicity.)
 - In fitting, we'd find the Gaussian distributions that best describe each class.
 - In prediction, we can take new images and classify them.
- Example #2 : Multinomial classification
 - Let's say the data are discrete.

- To be concrete, suppose each x_i is a document, i.e., a collection of observed words from a vocabulary.
- Suppose each class is described by a multinomial distribution. In this case, θ_k is a distribution over terms and we assume the words of each document are drawn independently from that distribution. That is,

$$p(x \mid \theta_k) = \prod_{j=1}^{\ell} \theta_{k, x_j}, \qquad (11)$$

where x_i is the *j*th word in document *x* (of length ℓ).

- Actually, we'll use the "selection" mechanism here too. This lets us see how the probability is only a function of the count of each word. (For this reason, models like this are often called "bag of words" models.)
- First, write down the previous equation in this form

$$p(x \mid \theta_k) = \prod_{j=1}^{\ell} \prod_{\nu=1}^{V} \theta_{k\nu}^{x_j^{\nu}}$$
(12)

- How many times does θ_{kv} appear in this product? The number of times *v* appears in *x*. This gives our final expression for the probability of document *x*,

$$p(x \mid \theta_k) = \prod_{\nu=1}^{V} \theta_{k\nu}^{n_{\nu}(x)}, \qquad (13)$$

where $n_v(x)$ is the number of times term v occurred in x.

- This is sometimes called a "Naive Bayes" classifier. (It's a silly name: Most models are naive and there isn't much Bayesian about this one.)

Prediction

- In prediction we are given the parameters, $\theta_{1:K}$ and π , and an unlabeled data point *x*. We want to predict the label for *x*.
- We use Bayes rule to compute the posterior distribution of the label

$$P(C|x) \propto P(x|C)P(C). \tag{14}$$

This is proportional because the denominator p(x) is constant with respect to *C*.

• For each possible label,

$$p(c|x) \propto p(x|\theta_c)\pi_c \tag{15}$$

- The precise form of the first term depends on the class-conditional data model.
- In the Gaussian case (with fixed variance) notice that

$$p(c|x) \propto \left(\frac{1}{2\sigma^2} (x - \mu_c)^2\right) \pi_c.$$
(16)

Everything else is constant with respect to the class label *c*.

- This equation says that we look at the squared difference between *x* and each class, weighted by the prior probability of that class.
- In the multinomial case, for each class we consider the probability that *x* was "generated" by its parameter, weighted again by the prior probability of each class,

$$p(c|x) \propto \left(\prod_{\nu=1}^{V} \theta_{c\nu}^{n_{\nu}(x)}\right) \pi_{c}$$
(17)

• In practice, we can take the label of maximum posterior probability. Or we can compute the probabilies and report a distribution.

Fitting

- Now we turn to the problem of finding parameters given data. We will find maximum likelihood estimates of the class conditional distributions θ_{1:K} and the class proportions π.
- The labeled data set is $\{x_i, c_i\}_{i=1}^n$. E.g.,
 - Labeled images
 - Labeled documents
 - Labeled genes
 - Labeled songs
- Taking the log of the product of joint distributions $p(c_i)p(x_i | c_i)$, the log likelihood is

$$\mathscr{L}(\pi, \theta_{1:K}) = \sum_{i=1}^{n} \sum_{j=1}^{k} c_{i}^{j} \log \pi^{j} + c_{i}^{j} \log p(x_{i} | \theta_{j})$$
(18)

- Finding the MLEs of π and $\theta_{1:K}$ decomposes into K + 1 MLE problems.
- First, the MLE of the class proportions is

$$\hat{\pi} = \arg\max_{\pi} \sum_{i=1}^{n} \sum_{j=1}^{k} c_i^j \log \pi^j$$
(19)

This is simply the empirical proportion of each class

$$\hat{\pi}^j = \frac{\sum_{i=1}^n c_i^j}{n},\tag{20}$$

where the numerator is the number of times we saw class *j*.

• The MLE of each class conditional parameter is

$$\hat{\theta}_j = \arg\max_{\theta_j} \sum_{i=1}^n c_i^k \log p(x_i | \theta_j).$$
(21)

Notice that

- Only the points labeled with class *j* play a role in this objective function.
- This is like taking a simple MLE of those points, drawn IID from θ_i .
- Operationally, take each class and compute the MLE of its parameter from the data assigned to it. In a Bernoulli case, compute the probability of heads. In a Guassian case, compute the empirical mean and variance.

Example: Simple Gaussian classification

- The data are average RGB values for images.
 - Whole images are summarized in a single color
- Draw the graphical model
 - Write the joint
 - Write the MLEs
 - Write the posterior given a new image
- Show the demo

Example: Multinomial classification

- Given labeled data, we can fit a model and classify new data points.
- This strategy is common in classifying documents. To review:
 - x_i is a collection of word counts.
 - x_i^u is the number of times word *u* occurred.
 - The class conditional parameter is a point on the term simplex,

$$\theta_{ku} > 0 \tag{22}$$

$$\sum_{u=1}^{\nu} \theta_{ku} = 1. \tag{23}$$

- The class conditional probability is

$$p(x_i^{\nu} | \theta_k) = \prod_{u=1}^{\nu} \theta_{ku}^{x_{iu}}$$
(24)

- This assumes that the collection of words came IID from θ_k . (You can confirm this.)
- To complete the algorithm, we only need the MLE of θ_i from the collection of documents.
- This is the proportion of times that each word occurred in each class *j* document:

$$\hat{\theta}_{j}^{u} = \frac{\sum_{i=1}^{n} c_{i}^{j} x_{i}^{u}}{\sum_{i=1}^{n} c_{i}^{j} \sum_{w=1}^{v} x_{i}^{w}}.$$
(25)

- Numerator: The number of times word *u* occurred in documents of class *j*
- Denominator: The number of words that occurred in documents of class j
- What happens if a test document contains a word that we never saw in training?
- In text models, we often *smooth* the parameter estimates.
 - The simplest smoother is to add a "pseudocount" to each word (such as one) before computing the MLE.
 - What does this do to the probabilities of frequent words? rare words?
 - Smoothing gets more complicated than that.
- It has interesting connections to

- Bayesian statistics: it can be construed as assuming the distribution came from a prior and then computing the posterior expectation of that distribution
- WWII history: A. Turing and I. J. Good developed smoothing to break the Nazi code.
- TODO: Implementing
 - The log trick