

Lecture 20: Counting and Sampling Problems

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Today's topic of counting and sampling problems is motivated by computational problems involving multivariate statistics and estimation, which arise in many fields. For instance, we may have a probability density function $\phi(x)$ where $x \in \mathbb{R}^n$. Then we may want to compute *moments* or other parameters of the distribution, e.g. $\int x^3 \phi(x) dx$. Or, we may have a model for how links develop faults in a network, and we seek to compute the probability that two nodes i, j stay connected under this model. This is a complicated probability calculation.

In general, such problems can be intractable (eg, NP-hard). The simple-looking problem of integrating a multivariate function is NP-hard in the worst case, even when we have an explicit expression for the function $f(x_1, x_2, \dots, x_n)$ that allows f to be computed in polynomial (in n) time.

$$\int_{x_1=0}^1 \int_{x_2=0}^1 \cdots \int_{x_n=0}^1 f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n.$$

In fact even approximating such integrals can be NP-hard, as shown by Koutis (2003).

Valiant (1979) showed that the computational heart of such problems is combinatorial *counting problems*. The goal in such problems is to compute the size of a set S where we can test *membership* in S in polynomial time. The class of such problems is called $\#P$.

EXAMPLE 1 $\#SAT$ is the problem where, given a boolean formula φ , we have to compute the *number* of satisfying assignments to φ . Clearly it is NP-hard since if we can solve it, we can in particular solve the *decision* problem: *decide* if the number of satisfying assignments at least 1.

$\#CYCLE$ is the problem where, given a graph $G = (V, E)$, we have to compute the number of *cycles* in G . Here the decision problem ("is G acyclic?") is easily solvable using breadth first search. Nevertheless, the counting problem turns out to be NP-hard.

$\#SPANNINGTREE$ is the problem where, given a graph $G = (V, E)$, we have to compute the number of *spanning trees* in G . This is known to be solvable using a simple determinant computation (Kirchoff's matrix-tree theorem) since the 19th century.

Valiant's class $\#P$ captures most interesting counting problems. Many of these are NP-hard, but not all. You can learn more about them in *COS 522: Computational Complexity*, usually taught in the spring semester. \square

It is easy to see that the above integration problem can be reduced to a counting problem with some loss of precision. First, recall that integration basically involves summation: we appropriately discretize the space and then take the sum of the integrand values (assuming in each cell of space the integrand doesn't vary much). Thus the integration reduces to some sum of the form

$$\sum_{x_1 \in [N], x_2 \in [N], \dots, x_n \in [N]} g(x_1, x_2, \dots, x_n),$$

where $[N]$ denotes the set of integers in $0, 1, \dots, N$. Now assuming $g(\cdot) \geq 0$ this is easily estimated using sizes of the following sets:

$$\{(x, c) : x \in [N]^n; c \leq g(x) \leq c + \epsilon\}.$$

Note if g is computable in polynomial time then we can test membership in this set in polynomial time given (x, c, ϵ) so we've shown that integration is a $\#P$ problem.

We will also be interested in *sampling* a random element of a set S . In fact, this will turn out to be intimately related to the problem of counting.

0.1 Counting vs Sampling

We say that an algorithm is an *approximation scheme* for a counting problem if for every $\epsilon > 0$ it can output an estimate of the size of the set that is correct within a multiplicative factor $(1 + \epsilon)$. We say it is a *randomized fully polynomial approximation scheme* (FPRAS) if it is randomized and it runs in $\text{poly}(n, 1/\epsilon, \log 1/\delta)$ time and has probability at least $(1 - \delta)$ of outputting such an answer. We will assume $\delta < 1/\text{poly}(n)$ so we can ignore the probability of outputting an incorrect answer.

An *fully polynomial-time approximate sampler* for S is one that runs in $\text{poly}(n, 1/\epsilon, \log 1/\delta)$ and outputs a sample $u \in S$ such that $\sum_{u \in S} \left| \Pr[u \text{ is output}] - \frac{1}{|S|} \right| \leq \epsilon$.

THEOREM 1 (JERRUM, VALIANT, VAZIRANI 1986)

For “nicely behaved” counting problems (the technical term is “downward self-reducible”) sampling in the above sense is equivalent to counting (i.e., a algorithm for one task can be converted into one for the other).

PROOF: For concreteness, let's prove this for the problem of counting the number of satisfying assignments to a boolean formula. Let $\#\varphi$ denote the number of satisfying assignments to formula φ .

Sampling \Rightarrow Approximate counting: Suppose we have an algorithm that is an approximate sampler for the set of satisfying assignments for any formula. For now assume it is an exact sampler instead of approximate. Take m samples from it and let p_0 be the fraction that have a 0 in the first bit x_1 , and p_1 be the fraction that have a 1. Assume $p_0 \geq 1/2$. Then the estimate of p_0 is correct up to factor $(1 + 1/\sqrt{m})$ by Chernoff bounds. But denoting by $\varphi|_{x_1=0}$ the formula obtained from φ by fixing x_1 to 0, we have

$$p_0 = \frac{\#\varphi|_{x_1=0}}{\#\varphi}.$$

Since we have a good estimate of p_0 , to get a good estimate of $\#\varphi$ it suffices to have a good estimate of $\#\varphi|_{x_1=0}$. So produce the formula $\varphi|_{x_1=0}$ obtained from φ by fixing x_1 to 0, then use the same algorithm recursively on this smaller formula to estimate N_0 , the value of $\#\varphi|_{x_1=0}$. Then output N_0/p_0 as your estimate of $\#\varphi$. (Base case $n = 1$ can be solved exactly of course.)

Thus if Err_n is the error in the estimate for formulae with n variables, this satisfies

$$\text{Err}_n \leq (1 + 1/\sqrt{m})\text{Err}_{n-1},$$

which solves to $\text{Err}_n \leq (1 + 1/\sqrt{m})^n$. By picking $m \gg n^2/\epsilon^2$ this error can be made less than $1 + \epsilon$. It is easily checked that if the sampler is not exact but only approximate, the algorithm works essentially unchanged, except the sampling error also enters the expression for the error in estimating p_0 .

Approximate counting \Rightarrow Sampling: This involves reversing the above reasoning. Given an approximate counting algorithm we are trying to generate a random satisfying assignment. First use the counting algorithm to approximate $\#\varphi|_{x_1=0}$ and $\#\varphi$ and take the ratio to get a good estimate of p_0 , the fraction of assignments that have 0 in the first bit. (If p_0 is too small, then we have a good estimate of $p_1 = 1 - p_0$.) Now toss a coin with $\text{Pr}[\text{heads}] = p_0$. If it comes up heads, output 0 as the first bit of the assignment and then recursively use the same algorithm on $\varphi|_{x_1=0}$ to generate the remaining $n - 1$ bits. If it comes up tails, output 1 as the first bit of the assignment and then recursively use the same algorithm on $\varphi|_{x_1=1}$ to generate the remaining $n - 1$ bits.

Note that the quality ϵ of the approximation suffers a bit in going between counting and sampling. \square

0.1.1 Monte Carlo method

The classical method to do counting via sampling is the *Monte Carlo* method. A simple example is the ancient method to estimate the area of a circle of unit radius. Draw the circle in a square of side 2. Now throw darts at the square and measure the fraction that fall in the circle. Multiply that fraction by 4 to get the area of the circle.

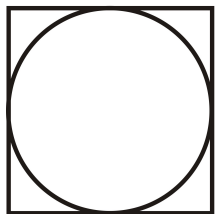


Figure 1: Monte Carlo (dart throwing) method to estimate the area of a circle. The fraction of darts that fall inside the disk is $\pi/4$.

Now replace “circle” with any set S and “square” with any set Ω that contains S and can be sampled in polynomial time. Then just take many samples from Ω and just observe the fraction that are in S . This is an estimate for $|S|$. The problem with this method is that usually the obvious Ω is much bigger than S , and we need $|\Omega|/|S|$ samples to get any that lie in S . (For instance the obvious Ω for computing $\#\varphi$ is the set of all possible assignments, which may be exponentially bigger.)

0.2 Dyer's algorithm for counting solutions to KNAPSACK

The Knapsack problem models the problem faced by a kid who is given a knapsack and told to buy any number of toys that fit in the knapsack. The problem is that not all toys give him the same happiness, so he has to trade off the happiness received from each toy with its size; toys with high happiness/size ratio are preferred. Turns out this problem is NP-hard if the numbers are given in binary. We are interested in a counting version of the problem that uses just the sizes.

DEFINITION 1 *Given n weights w_1, w_2, \dots, w_n and a target weight W , a feasible solution to the knapsack problem is a subset T such that $\sum_{i \in T} w_i \leq W$.*

We wish to approximately count the number of feasible solutions. This had been the subject of some very technical papers, until M. Dyer gave a very elementary solution in 2003.

First, we note that the counting problem can be solved *exactly* in $O(nW)$ time, though of course this is not polynomial since W is given to us in binary, i.e. using $\log W$ bits. The idea is dynamic programming. Let $\text{Count}(i, U)$ denote the number of feasible solutions involving only the first i numbers, and whose total weight is at most U . The dynamic programming follows by observing that there are two types of solutions: those that involve the i th element, and those that don't. Thus

$$\text{Count}(i, U) = \begin{cases} \text{Count}(i-1, U - w_i) + \text{Count}(i-1, U) & \\ 1 & \text{if } i = 1 \text{ and } w_1 \leq U \\ 0 & \text{if } i = 1 \text{ and } w_1 > U \end{cases}$$

Denoting by S the set of feasible solutions, $|S| = \text{Count}(n, W)$. But as observed, computing this exactly is computationally expensive and not polynomial-time. Dyer's next idea is to find a set Ω containing S but at most n times bigger. This set Ω can be exactly counted as well as sampled from. So then by the Monte Carlo method we can estimate the size of S in polynomial time by drawing samples from Ω .

Ω is simply the set of solutions to a Knapsack instance in which the weights have been rounded to lie in $[0, n^2]$. Specifically, let $w'_i = \lfloor \frac{w_i n^2}{W} \rfloor$ and $W' = n^2$. Then Ω is the set of feasible solutions to this modified knapsack problem.

CLAIM 1: $S \subseteq \Omega$. (Consequently, $|S| \leq |\Omega|$.)

This follows since if $T \in S$ is a feasible solution for the original problem, then $\sum_i w'_i \leq \sum_i w_i n^2 / W \leq n^2$, and so T is a feasible solution for the rounded problem.

CLAIM 2: $|\Omega| \leq n|S|$.

To prove this we give a mapping g from Ω to S that is at most n -to-1.

$$g(T') = \begin{cases} = T' & \text{if } T' \in S \\ = T' \setminus \{j\} & \text{(else) where } j = \text{index of element in } T' \text{ with highest value of } w'_j \end{cases}$$

In the second case note that this element j satisfies $w_j > W/n$ which implies $w'_j \geq n$.

Clearly, g is at most n -to-1 since a set T in S can have at most n pre-images under g . Now let's verify that $T = g(T')$ lies in S .

$$\begin{aligned} \sum_{i \in T} w_i &\leq \sum_{i \in T} \frac{W}{n^2} (w'_i + 1) \\ &\leq \frac{W}{n^2} \times (W' - w'_j + n - 1) \\ &\leq W \quad (\text{since } W' = n^2 \text{ and } w'_j \geq n) \end{aligned}$$

which implies $T \in S$. \square

Sampling algorithm for Ω To sample from Ω , use our earlier equivalence of approximate counting and sampling. That algorithm needs an approximate count not only for $|\Omega|$ but also for the subset of Ω that contain the first element. This is another knapsack problem and can thus be solved by Dyer's dynamic programming. And same is true for instances obtained in the recursion.

Bibliography

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