Simulation wrap-up

Last time

- Time-driven, event-driven
- "Simulation" from differential equations
- Cellular automata, microsimulation, agent-based simulation
 - see e.g.
 <u>http://www.microsimulation.org/IMA/What%20is%20</u>
 <u>microsimulation.htm</u>
- Example applications: SIR disease model, population genetics

Simulation: Pros and Cons

• Pros:

- Building model can be easy (easier) than other approaches
- Outcomes can be easy to understand
- Cheap, safe
- Good for comparisons
- Cons:
 - Hard to debug
 - No guarantee of optimality
 - Hard to establish validity
 - Can't produce absolute numbers

Simulation: Important Considerations

- Are outcomes statistically significant? (Need many simulation runs to assess this)
- What should initial state be?
- How long should the simulation run?
- Is the model realistic?
- How sensitive is the model to parameters, initial conditions?

Statistics Overview

Descriptive statistics



Random Variables

- A random variable is any "probabilistic outcome"
 - e.g., a coin flip, height of someone randomly chosen from a population
- A R.V. takes on a value in a sample space
 - space can be discrete, e.g., {H, T}
 - or continuous, e.g. height in (0, infinity)
- R.V. denoted with capital letter (X), a realization with lowercase letter (x)
 - e.g., X is a coin flip, x is the value (H or T) of that coin flip

Probability Mass Function

• Describes probability for a discrete R.V.



Probability Density Function

• Describes probability for a continuous R.V.



[Population] Mean of a Random Variable

- aka expected value, first moment
- for discrete RV: $E[X] = \mu = \sum x_i p_i$

(requires that
$$\sum_{i}^{i} p_{i} = 1$$
)

• for continuous RV: $E[X] = \mu = \int_{-\infty}^{\infty} x p(x) dx$

(requires that
$$\int_{-\infty}^{\infty} p(x) dx = 1$$
)

[Population] Variance

$$\sigma^{2} = \mathbb{E}\left[(X - \mu)^{2}\right]$$
$$= \mathbb{E}\left[X^{2} - 2X\mu + \mu^{2}\right]$$
$$= \mathbb{E}\left[X^{2}\right] - \mu^{2}$$
$$= \mathbb{E}\left[X^{2}\right] - \left(\mathbb{E}\left[X\right]\right)^{2}$$
• for discrete RV:
$$\sigma^{2} = \sum_{i} p_{i}(x_{i} - \mu)^{2}$$

• for continuous RV:

$$\sigma^2 = \int (x - \mu)^2 p(x) dx$$

Sample mean and sample variance

- Suppose we have N independent observations of X: x₁, x₂, ...x_N
- Sample mean:

Unbiased:

$$\frac{1}{N} \sum_{i=1}^{N} x_i = \overline{x}$$

 $\mathrm{E}[\bar{x}] = \mu$

 $E[s^2] = \sigma^2$

• Sample variance:

$$\frac{1}{N-1} \sum_{i=1}^{N} (x_i - \bar{x})^2 = s^2$$

1/(N-1) and the sample variance

• The N differences $x_i - \overline{x}$ are not independent:

$$\sum (x_i - \bar{x}) = 0$$

 If you know N-1 of these values, you can deduce the last one

• Could treat sample as population and compute population variance: $1 \sum_{n=2}^{N} \frac{1}{2} \sum_{n=2}^{N} \frac{$

$$\frac{1}{N}\sum_{i=1}^{N}(x_i-\overline{x})^2$$

 BUT this underestimates true population variance (especially bad if sample is small)

Sample variance using 1/(N-1) is unbiased

$$E\left[s^{2}\right] = E\left[\frac{1}{N-1}\sum_{i=1}^{N}(x_{i}-\bar{x})^{2}\right]$$
$$= \frac{1}{N-1}E\left[\sum_{i=1}^{N}x_{i}^{2}-N\bar{x}^{2}\right]$$
$$= \frac{1}{N-1}\left[N\left(\sigma^{2}+\mu^{2}\right)-N\left(\frac{\sigma^{2}}{N}+\mu^{2}\right)\right]$$
$$= \sigma^{2}$$

Computing Sample Variance

• Can compute as

$$s^{2} = \frac{1}{N-1} \sum_{i=1}^{N} (x_{i} - \bar{x})^{2}$$

• Prefer:

$$s^{2} = \frac{\left(\sum_{i=1}^{N} x_{i}^{2}\right) - N(\bar{x})^{2}}{N-1}$$

(one pass, fewer operations, more accurate)

The Gaussian Distribution





- sum of independent observations of random variables converges to Gaussian *(with some assumptions)
- in nature, events having variations resulting from many small, independent effects tend to have Gaussian distributions
 - demo: <u>http://www.mongrav.org/math/falling-balls-</u> probability.htm
 - e.g., measurement error
 - if effects are multiplicative, logarithm is often normally distributed

Central Limit Theorem

• Suppose we sample $x_1, x_2, ..., x_N$ from a distribution with mean μ and variance σ^2

• Let
$$\overline{x} = \frac{1}{N} \sum_{i=1}^{N} x_i$$

holds for *(almost) any parent distribution!

• then

$$z = \frac{\overline{x} - \mu}{\sigma / \sqrt{N}} \rightarrow N(0, 1)$$

• i.e., \overline{x} distributed normally with mean μ , variance σ^2/N

- 1. Family of normal distributions closed under linear transformations:
 - if $X \sim N(\mu, \sigma^2)$ then
 - $(aX + b) \sim N(a\mu+b, a^2\sigma^2)$
- 2. Linear combination of normals is also normal:

if $X_1 \sim N(\mu_1, \sigma_1^2)$ and $X_2 \sim N(\mu_2, \sigma_2^2)$ then a $X_1 + bX_2 \sim N(a\mu_1 + b\mu_2, a^2\sigma_1^2 + b^2\sigma_2^2)$

- 3. Of all distributions with mean and variance, normal has **maximum entropy**
 - Information theory: Entropy like "uninformativeness"
 - Principle of maximum entropy: choose to represent the world with as uninformative a distribution as possible, subject to "testable information"
- If we know x is in [a, b], then uniform distribution on [a, b] has least entropy
- If we know distribution has mean μ , variance σ^2 , normal distribution N(μ , σ^2) has least entropy

- 4. If errors are normally distributed, a least-squares fit yields the maximum likelihood estimator
- Finding least-squares x st $Ax \approx b$ finds the value of x that maximizes the likelihood of data A under some model

$$P(data|model) \propto \prod_{i=1}^{n} \exp\left[-\frac{1}{2}\left(\frac{y_i - y(x_i)}{\sigma_i}\right)^2\right] \Delta y$$

 $P(model|data) \propto P(data|model)P(model)$

5. Many derived random variables have analytically-known densities
e.g., sample mean, sample variance
6. Sample mean and variance of n identical independent samples are independent; sample mean is a normally-distributed random variable

$$\overline{X}_n \sim N(\mu, \sigma^2/n)$$

Distribution of Sample Variance (For Gaussian R.V. X)

$$s^{2} = \frac{1}{N-1} \sum_{i=1}^{N} (x_{i} - \overline{x})^{2}$$

define $U = \frac{(n-1)s^{2}}{\sigma^{2}}$

then U has a χ^2 distribution with (n-1) d.o.f.

$$p(x) = \left[2^{n/2} \Gamma\left(\frac{n}{2}\right)\right]^{-1} (x)^{\frac{n}{2}-1} e^{-x/2}, \quad x \ge 0$$
$$E[U] = n - 1, \quad Var[U] = 2(n - 1)$$

The Gamma Function

$$\Gamma(z) = \int_0^\infty t^{z-1} e^{-t} \,\mathrm{d}t \,.$$

$$\Gamma(n) = (n-1)!$$

The Chi-Squared Distribution



What if we don't know true variance?

- Sample mean is normally distributed R.V. $\overline{X}_n \sim N(\mu, \sigma^2/n)$
- Taking advantage of this presumes we know σ^2

•
$$\frac{\overline{x} - \mu}{s_n / \sqrt{n}}$$
 has a *t* distribution with (n-1) d.o.f.

[Student's] t-distribution



Forming a confidence interval

e.g., given that I observed a sample mean of _____, I'm 99% confident that the true mean lies between _____ and ____.

• Know that
$$\frac{x-\mu}{s_n/\sqrt{n}}$$
 has t distribution

• Choose q_1 , q_2 such that student t with (n-1) dof has 99% probability of lying between q_1 , q_2



Confidence interval for the mean

if P
$$\left\{ q_1 < \frac{\bar{x}_n - \mu}{s_n / \sqrt{n}} < q_2 \right\} = 0.99$$

then P $\left\{ \bar{x}_n - q_2 \frac{s_n}{\sqrt{n}} < \mu < \bar{x}_n - q_1 \frac{s_n}{\sqrt{n}} \right\} = 0.99$

Interpreting Simulation Outcomes

- How long will customers have to wait, on average?
 - e.g., for given # tellers, arrival rate, service time distribution, etc.



Interpreting Simulation Outcomes

- Simulate bank for N customers
- Let x_i be the wait time of customer i
- Is mean(x) a good estimate for µ?
- How to compute a 95% confidence interval for µ?
 Problem: x_i are not independent!

Replications

- Run simulation to get M observations
- Repeat simulation N times (different random numbers each time)
- Treat the sample mean of different runs as approximately uncorrelated

$$\overline{\mathbf{X}}_{\mathbf{i}} = \frac{1}{\mathbf{m}} \sum_{j=1}^{\mathbf{n}} \mathbf{x}_{ij} \qquad \overline{\overline{\mathbf{X}}}_{i} = \frac{1}{\mathbf{n}} \sum_{i=1}^{\mathbf{n}} \overline{\mathbf{X}}_{i}$$
$$s^{2} = \frac{1}{n-1} \sum_{i} (\overline{X}_{i} - \overline{\overline{X}})^{2}$$